

Lecture 23

Last time

- ▷ Markov Property
- ▷ consequences

Today

- ▷ Strong Markov Property
- ▷ consequences

Question: Imagine we start $B_0 = 0$, then what is the less likely point $t \in [0, 1]$ to be the last point in $[0, 1]$ $B_t = 0$? What's the distribution of $L = \sup\{t : B_t = 0\}$?

The answer to the first question is $1/2$!

Stopping times

We say that a filtration $\{\mathcal{F}_t\}$ is right continuous if

$$\bigcap_{s>t} \mathcal{F}_s = \mathcal{F}_t.$$

The reason we like right continuous filtrations is that they make infinitesimals into the future negligible.

Def: We say that a random variable

S in $[0, \infty]$ is a **stopping time** w.r.t. a filtration $\{\mathcal{F}_t\}$ if $\{S \leq t\} \in \mathcal{F}_t$ $\forall t$.

This is the same as $\{S \leq t\} \in \mathcal{F}_t$ if $\{\mathcal{F}_t\}$ is r.c. Why?

We need to understand what type of things are stopping times?

Q: Say that $G \subseteq \mathbb{R}$ is an open set or a closed set. Is $T_G = \inf\{t : B_t \in G\}$ a stopping time? **Yes!**

Theorem: If G is an open set, then T_G is a stopping time.

Proof: Since G is open and $t \rightarrow B_t$ is continuous, then

$$\{T < t\} = \bigcup_{\substack{q < t \\ q \in \mathbb{Q}}} \{B_q \in G\},$$

thus we conclude that $\{T < t\} \in \mathcal{F}_t$. \square

Theorem: Suppose that T_n is a sequence of stopping times. If either

$$T_n \downarrow T \quad \text{or} \quad T_n \uparrow T.$$

Then, T is a stopping time.

Proof: It suffices to note that

$$\{T < t\} = \bigcup_n \{T_n < t\} \quad \text{and} \quad \{T \leq t\} = \bigcap_n \{T_n \leq t\}$$

Theorem: If G is closed, then T_G is a stopping time.

Proof: Let $B(x, r) = \{y: |y - x| < r\}$, let

$G_n = \bigcup_{x \in K} B(x, \frac{1}{n})$ and let $T_n = \inf\{t \geq 0: B_t \in G_n\}$. Since G_n is open $\Rightarrow T_n$ is a stopping time.

Next we show that $T_n \uparrow T$. Notice

that by construction $T_n \leq T$ and

$T_n \uparrow t^*$ for some $t^* < \infty$. Since $B_{T_n} \in \overline{G_n} \forall n \Rightarrow B_{T_n} \rightarrow B_{t^*} \in G$ and so

$t^* \geq T \Rightarrow \lim T_n = T. \quad \square$

Strong Markov Property

We develop an analogue of the SMP. We now define the random shift operator. Given a nonnegative r.v. S in $[0, \infty]$, define

$$(\theta_S(\omega))(t) = \begin{cases} \omega(S(\omega) + t) & \text{on } \{S < \infty\} \\ \Delta & \text{on } \{S = \infty\} \end{cases}$$

↖ Extra symbol.

We also define the information known at time s :

$$\mathcal{F}_s = \{A: A \cap \{s \leq t\} \in \mathcal{F}_t \text{ for all } t \geq 0\}.$$

Proposition: If $S \leq T$ are stopping times
 $\Rightarrow \mathcal{F}_S \subseteq \mathcal{F}_T$.

Proposition: If $T_n \downarrow T$ are stopping times
 $\Rightarrow \mathcal{F}_T = \bigcap \mathcal{F}_{T_n}$.

Exercise: Prove these two facts!

Theorem (Strong Markov Property). Let $(s, \omega) \rightarrow Y_s(\omega)$ be bounded and $\mathbb{R} \times \Omega$ measurable. If S is a stopping time, then for all $x \in \mathbb{R}$

$$\mathbb{E}_x [Y_S \circ \theta_S \mid \mathcal{F}_S] = \mathbb{E}_{B_S} Y_S \text{ on } \{S < \infty\}.$$

Function $\varphi(x, t) = \mathbb{E}_x Y_t$
evaluated at $x = B_S$ and $t = S$.

The proof of this result is similar to the one we covered in Lecture 16 (albeit much more technical), see Theorem

7.3.9. in Durrett.

Reflection Principle

Let $a > 0$ and $T_a = \inf \{ t : B_t = a \}$

Theorem:

$$P_0(T_a < t) = 2 P_0(B_t \geq a).$$

Proof: We shall see that

$$P_0(T_a < t, B_t > a) = \frac{1}{2} P_0(T_a < t), \quad (\heartsuit)$$

which right away implies:

$$P_0(T_a < t) = 2 P_0(B_t \geq a)$$

↑ since $\{T_a < t\} \supseteq \{B_t > a\}$.

We focus on (\heartsuit) . We will use the SMP, define

$$Y_s(\omega) = \begin{cases} 1 & \text{if } s < t, \omega(t-s) > a, \\ 0 & \text{otherwise.} \end{cases}$$

If we let $s = \inf \{ s < t : B_s = a \}$ with $\inf \emptyset = \infty$, then

$$Y_s(\theta_s(\omega)) = \begin{cases} 1 & \text{if } s < t, B_t > a \\ 0 & \text{otherwise.} \end{cases}$$

So SMP gives

$$\rightarrow \mathbb{E}_0(Y_s \circ \theta_s | \mathcal{F}_s) = \mathbb{E}_{B_s}(Y_s) = \mathbb{E}_a(Y_s)$$

on $\{s < \infty\} = \{T_a < t\}$

$$= \frac{1}{2}$$

← Gaussian centered at a .

Taking expectations

$$\begin{aligned}
 P_0(T_a < t, B_t \geq a) &= \mathbb{E}_0 [(Y_s \circ \theta_s) \mathbb{1}_{Y_s < \infty}] \\
 &= \mathbb{E}_0 [\mathbb{E}_0 [Y_s \circ \theta_s | \mathcal{F}_s] \mathbb{1}_{Y_s < \infty}] \\
 &= \frac{1}{2} \mathbb{E}_0 [\mathbb{1}_{\{T_a < t\}}] \\
 &= \frac{1}{2} P(T_a < t). \quad \square
 \end{aligned}$$

But this means that we have a closed form for the dist. of T_a :

$$P_0(T_a \leq t) = 2 P_0(B_t \geq a) = \frac{2}{\sqrt{2\pi t}} \int_a^\infty \exp(-x^2/2t) dx.$$

Recall we were interested in $L = \inf \{t \leq 1 : B_t = 1\}$.

Then,

$$\begin{aligned}
 P_0(L \leq t) &\stackrel{\text{HW } \delta}{=} \int_{-\infty}^{-\infty} P_t(0, y) P_y(T_0 > 1-t) dy \\
 &= 2 \int_0^\infty (2\pi t)^{-1/2} \exp(-y^2/2t) \\
 &\quad \int_{1-s}^\infty (2\pi u)^{-1/2} \exp(-x^2/2t) dx
 \end{aligned}$$

∴ ← Change of variables and simplifications

$$= \frac{1}{\pi} \int_0^t (s(1-s))^{-1/2} ds = \frac{2}{\pi} \arcsin(\sqrt{s})$$

Thus, the density of L is $\frac{1}{\pi} (s(1-s))^{-1/2}$
which looks like

